

HW Solution 10 — Due: Nov 18, 8:59 AM

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Problem 1 (Randomly Phased Sinusoid). Suppose Θ is a uniform random variable on the interval $(0, 2\pi)$.

- (a) Consider another random variable X defined by

$$X = 5 \cos(7t + \Theta)$$

where t is some constant. Find $\mathbb{E}[X]$.

- (b) Consider another random variable Y defined by

$$Y = 5 \cos(7t_1 + \Theta) \times 5 \cos(7t_2 + \Theta)$$

where t_1 and t_2 are some constants. Find $\mathbb{E}[Y]$.

Solution: First, because Θ is a uniform random variable on the interval $(0, 2\pi)$, we know that $f_{\Theta}(\theta) = \frac{1}{2\pi} \mathbf{1}_{(0, 2\pi)}(\theta)$. Therefore, for “any” function g , we have

$$\mathbb{E}[g(\Theta)] = \int_{-\infty}^{\infty} g(\theta) f_{\Theta}(\theta) d\theta.$$

- (a) X is a function of Θ . $\mathbb{E}[X] = 5\mathbb{E}[\cos(7t + \Theta)] = 5 \int_0^{2\pi} \frac{1}{2\pi} \cos(7t + \theta) d\theta$. Now, we know that integration over a cycle of a sinusoid gives 0. So, $\mathbb{E}[X] = \boxed{0}$.

- (b) Y is another function of Θ .

$$\begin{aligned} \mathbb{E}[Y] &= \mathbb{E}[5 \cos(7t_1 + \Theta) \times 5 \cos(7t_2 + \Theta)] = \int_0^{2\pi} \frac{1}{2\pi} 5 \cos(7t_1 + \theta) \times 5 \cos(7t_2 + \theta) d\theta \\ &= \frac{25}{2\pi} \int_0^{2\pi} \cos(7t_1 + \theta) \times \cos(7t_2 + \theta) d\theta. \end{aligned}$$

Recall¹ the cosine identity

$$\cos(a) \times \cos(b) = \frac{1}{2} (\cos(a + b) + \cos(a - b)).$$

Therefore,

$$\begin{aligned} \mathbb{E}Y &= \frac{25}{4\pi} \int_0^{2\pi} \cos(14t + 2\theta) + \cos(7(t_1 - t_2)) d\theta \\ &= \frac{25}{4\pi} \left(\int_0^{2\pi} \cos(14t + 2\theta) d\theta + \int_0^{2\pi} \cos(7(t_1 - t_2)) d\theta \right). \end{aligned}$$

The first integral gives 0 because it is an integration over two period of a sinusoid. The integrand in the second integral is a constant. So,

$$\mathbb{E}Y = \frac{25}{4\pi} \cos(7(t_1 - t_2)) \int_0^{2\pi} d\theta = \frac{25}{4\pi} \cos(7(t_1 - t_2)) 2\pi = \boxed{\frac{25}{2} \cos(7(t_1 - t_2))}.$$

Problem 2 (Yates and Goodman, 2005, Q3.4.5). X is a continuous uniform RV on the interval $(-5, 5)$.

- What is its pdf $f_X(x)$?
- What is its cdf $F_X(x)$?
- What is $\mathbb{E}[X]$?
- What is $\mathbb{E}[X^5]$?
- What is $\mathbb{E}[e^X]$?

Solution: For a uniform random variable X on the interval (a, b) , we know that

$$f_X(x) = \begin{cases} 0, & x < a \text{ or } x > b, \\ \frac{1}{b-a}, & a \leq x \leq b \end{cases}$$

¹This identity could be derived easily via the Euler's identity:

$$\begin{aligned} \cos(a) \times \cos(b) &= \frac{e^{ja} + e^{-ja}}{2} \times \frac{e^{jb} + e^{-jb}}{2} = \frac{1}{4} (e^{ja}e^{jb} + e^{-ja}e^{jb} + e^{ja}e^{-jb} + e^{-ja}e^{-jb}) \\ &= \frac{1}{2} \left(\frac{e^{ja}e^{jb} + e^{-ja}e^{-jb}}{2} + \frac{e^{-ja}e^{jb} + e^{ja}e^{-jb}}{2} \right) \\ &= \frac{1}{2} (\cos(a + b) + \cos(a - b)). \end{aligned}$$

and

$$F_X(x) = \begin{cases} 0, & x < a \text{ or } x > b, \\ \frac{x-a}{b-a}, & a \leq x \leq b. \end{cases}$$

In this problem, we have $a = -5$ and $b = 5$.

$$(a) f_X(x) = \begin{cases} 0, & x < -5 \text{ or } x > 5, \\ \frac{1}{10}, & -5 \leq x \leq 5 \end{cases}$$

$$(b) F_X(x) = \begin{cases} 0, & x < a \text{ or } x > b, \\ \frac{x+5}{10}, & a \leq x \leq b. \end{cases}$$

$$(c) \mathbb{E}X = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-5}^5 x \times \frac{1}{10} dx = \frac{1}{10} \frac{x^2}{2} \Big|_{-5}^5 = \frac{1}{20} (5^2 - (-5)^2) = \boxed{0}.$$

In general,

$$\mathbb{E}X = \int_a^b x \frac{1}{b-a} dx = \frac{1}{b-a} \int_a^b x dx = \frac{1}{b-a} \frac{x^2}{2} \Big|_a^b = \frac{1}{b-a} \frac{b^2 - a^2}{2} = \frac{a+b}{2}.$$

With $a = -5$ and $b = 5$, we have $\mathbb{E}X = \boxed{0}$.

$$(d) \mathbb{E}[X^5] = \int_{-\infty}^{\infty} x^5 f_X(x) dx = \int_{-5}^5 x^5 \times \frac{1}{10} dx = \frac{1}{10} \frac{x^6}{6} \Big|_{-5}^5 = \frac{1}{60} (5^6 - (-5)^6) = \boxed{0}.$$

In general,

$$\mathbb{E}[X^5] = \int_a^b x^5 \frac{1}{b-a} dx = \frac{1}{b-a} \int_a^b x^5 dx = \frac{1}{b-a} \frac{x^6}{6} \Big|_a^b = \frac{1}{b-a} \frac{b^6 - a^6}{6}.$$

With $a = -5$ and $b = 5$, we have $\mathbb{E}[X^5] = \boxed{0}$.

(e) In general,

$$\mathbb{E}[e^X] = \int_a^b e^x \frac{1}{b-a} dx = \frac{1}{b-a} \int_a^b e^x dx = \frac{1}{b-a} e^x \Big|_a^b = \frac{e^b - e^a}{b-a}.$$

With $a = -5$ and $b = 5$, we have $\mathbb{E}[e^X] = \boxed{\frac{e^5 - e^{-5}}{10}} \approx 14.84$.

Problem 3. A random variable X is a Gaussian random variable if its pdf is given by

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-m}{\sigma}\right)^2},$$

for some constant m and positive number σ . Furthermore, when a Gaussian random variable has $m = 0$ and $\sigma = 1$, we say that it is a standard Gaussian random variable. There is no closed-form expression for the cdf of the standard Gaussian random variable. The cdf itself is denoted by Φ and its values (or its complementary values $Q(\cdot) = 1 - \Phi(\cdot)$) are traditionally provided by a table.

Suppose Z is a standard Gaussian random variable.

(a) Use the Φ table to find the following probabilities:

- (i) $P[Z < 1.52]$
- (ii) $P[Z < -1.52]$
- (iii) $P[Z > 1.52]$
- (iv) $P[Z > -1.52]$
- (v) $P[-1.36 < Z < 1.52]$

(b) Use the Φ table to find the value of c that satisfies each of the following relation.

- (i) $P[Z > c] = 0.14$
- (ii) $P[-c < Z < c] = 0.95$

Solution:

(a)

- (i) $P[Z < 1.52] = \Phi(1.52) = \boxed{0.9357}$.
- (ii) $P[Z < -1.52] = \Phi(-1.52) = 1 - \Phi(1.52) = 1 - 0.9357 = \boxed{0.0643}$.
- (iii) $P[Z > 1.52] = 1 - P[Z < 1.52] = 1 - \Phi(1.52) = 1 - 0.9357 = \boxed{0.0643}$.
- (iv) It is straightforward to see that the area of $P[Z > -1.52]$ is the same as $P[Z < 1.52] = \Phi(1.52)$. So, $P[Z > -1.52] = \boxed{0.9357}$.
Alternatively, $P[Z > -1.52] = 1 - P[Z \leq -1.52] = 1 - \Phi(-1.52) = 1 - (1 - \Phi(1.52)) = \Phi(1.52)$.
- (v) $P[-1.36 < Z < 1.52] = \Phi(1.52) - \Phi(-1.36) = \Phi(1.52) - (1 - \Phi(1.36)) = \Phi(1.52) + \Phi(1.36) - 1 = 0.9357 + 0.9131 - 1 = \boxed{0.8488}$.

(b)

- (i) $P[Z > c] = 1 - P[Z \leq c] = 1 - \Phi(c)$. So, we need $1 - \Phi(c) = 0.14$ or $\Phi(c) = 1 - 0.14 = 0.86$. In the Φ table, we do not have exactly 0.86, but we have 0.8599 and 0.8621. Because 0.86 is closer to 0.8599, we answer the value of c whose $\phi(c) = 0.8599$. Therefore, $c \approx \boxed{1.08}$.
- (ii) $P[-c < Z < c] = \Phi(c) - \Phi(-c) = \Phi(c) - (1 - \Phi(c)) = 2\Phi(c) - 1$. So, we need $2\Phi(c) - 1 = 0.95$ or $\Phi(c) = 0.975$. From the Φ table, we have $c \approx \boxed{1.96}$.

Problem 4. The peak temperature T , as measured in degrees Fahrenheit, on a July day in New Jersey is a $\mathcal{N}(85, 100)$ random variable.

Remark: Do not forget that, for our class, the second parameter in $\mathcal{N}(\cdot, \cdot)$ is the variance (not the standard deviation).

- (a) Express the cdf of T in terms of the Φ function.
- (b) Express each of the following probabilities in terms of the Φ function(s). Make sure that the arguments of the Φ functions are positive. (Positivity is required so that we can directly use the Φ/Q tables to evaluate the probabilities.)
- (i) $P[T > 100]$
- (ii) $P[T < 60]$
- (iii) $P[70 \leq T \leq 100]$
- (c) Express each of the probabilities in part (b) in terms of the Q function(s). Again, make sure that the arguments of the Q functions are positive.
- (d) Evaluate each of the probabilities in part (b) using the Φ/Q tables.
- (e) Observe that the Φ table (“Table 4” from the lecture) stops at $z = 2.99$ and the Q table (“Table 5” from the lecture) starts at $z = 3.00$. Why is it better to give a table for $Q(z)$ instead of $\Phi(z)$ when z is large?

Solution:

- (a) Recall that when $X \sim \mathcal{N}(m, \sigma^2)$, $F_X(x) = \Phi\left(\frac{x-m}{\sigma}\right)$. Here, $T \sim \mathcal{N}(85, 10^2)$. Therefore,

$$F_T(t) = \Phi\left(\frac{t - 85}{10}\right).$$

(b)

(i) $P[T > 100] = 1 - P[T \leq 100] = 1 - F_T(100) = 1 - \Phi\left(\frac{100-85}{10}\right) = 1 - \Phi(1.5)$

(ii) $P[T < 60] = P[T \leq 60]$ because T is a continuous random variable and hence $P[T = 60] = 0$. Now, $P[T \leq 60] = F_T(60) = \Phi\left(\frac{60-85}{10}\right) = \Phi(-2.5) = 1 - \Phi(2.5)$. Note that, for the last equality, we use the fact that $\Phi(-z) = 1 - \Phi(z)$.

(iii)

$$\begin{aligned}
 P[70 \leq T \leq 100] &= F_T(100) - F_T(70) = \Phi\left(\frac{100-85}{10}\right) - \Phi\left(\frac{70-85}{10}\right) \\
 &= \Phi(1.5) - \Phi(-1.5) = \Phi(1.5) - (1 - \Phi(1.5)) = 2\Phi(1.5) - 1.
 \end{aligned}$$

(c) In this question, we use the fact that $Q(x) = 1 - \Phi(x)$.

(i) $1 - \Phi(1.5) = Q(1.5)$.

(ii) $1 - \Phi(2.5) = Q(2.5)$.

(iii) $2\Phi(1.5) - 1 = 2(1 - Q(1.5)) - 1 = 2 - 2Q(1.5) - 1 = 1 - 2Q(1.5)$.

(d)

(i) $1 - \Phi(1.5) = 1 - 0.9332 = 0.0668$.

(ii) $1 - \Phi(2.5) = 1 - 0.99379 = 0.0062$.

(iii) $2\Phi(1.5) - 1 = 2(0.9332) - 1 = 0.8664$.

(e) When z is large, $\Phi(z)$ will start with 0.999... The first few significant digits will all be the same and hence not quite useful to be there.

Extra Questions

Here are some optional questions for those who want more practice.

Problem 5. Let X be a uniform random variable on the interval $[0, 1]$. Set

$$A = \left[0, \frac{1}{2}\right), \quad B = \left[0, \frac{1}{4}\right) \cup \left[\frac{1}{2}, \frac{3}{4}\right), \quad \text{and} \quad C = \left[0, \frac{1}{8}\right) \cup \left[\frac{1}{4}, \frac{3}{8}\right) \cup \left[\frac{1}{2}, \frac{5}{8}\right) \cup \left[\frac{3}{4}, \frac{7}{8}\right).$$

Are the events $[X \in A]$, $[X \in B]$, and $[X \in C]$ independent?

Solution: Note that

$$P[X \in A] = \int_0^{\frac{1}{2}} dx = \frac{1}{2},$$

$$P[X \in B] = \int_0^{\frac{1}{4}} dx + \int_{\frac{1}{2}}^{\frac{3}{4}} dx = \frac{1}{2}, \text{ and}$$

$$P[X \in C] = \int_0^{\frac{1}{8}} dx + \int_{\frac{1}{4}}^{\frac{3}{8}} dx + \int_{\frac{1}{2}}^{\frac{5}{8}} dx + \int_{\frac{3}{4}}^{\frac{7}{8}} dx = \frac{1}{2}.$$

Now, for pairs of events, we have

$$P([X \in A] \cap [X \in B]) = \int_0^{\frac{1}{4}} dx = \frac{1}{4} = P[X \in A] \times P[X \in B], \quad (10.1)$$

$$P([X \in A] \cap [X \in C]) = \int_0^{\frac{1}{8}} dx + \int_{\frac{1}{4}}^{\frac{3}{8}} dx = \frac{1}{4} = P[X \in A] \times P[X \in C], \text{ and} \quad (10.2)$$

$$P([X \in B] \cap [X \in C]) = \int_0^{\frac{1}{8}} dx + \int_{\frac{1}{2}}^{\frac{5}{8}} dx = \frac{1}{4} = P[X \in B] \times P[X \in C]. \quad (10.3)$$

Finally,

$$P([X \in A] \cap [X \in B] \cap [X \in C]) = \int_0^{\frac{1}{8}} dx = \frac{1}{8} = P[X \in A] P[X \in B] P[X \in C]. \quad (10.4)$$

From (10.1), (10.2), (10.3) and (10.4), we can conclude that the events $[X \in A]$, $[X \in B]$, and $[X \in C]$ are independent.

Problem 6 (Q3.5.6). Solve this question using the Φ/Q table.

A professor pays 25 cents for each blackboard error made in lecture to the student who points out the error. In a career of n years filled with blackboard errors, the total amount in dollars paid can be approximated by a Gaussian random variable Y_n with expected value $40n$ and variance $100n$.

- (a) What is the probability that Y_{20} exceeds 1000?
- (b) How many years n must the professor teach in order that $P[Y_n > 1000] > 0.99$?

Solution: We are given² that $Y_n \sim \mathcal{N}(40n, 100n)$. Recall that when $X \sim \mathcal{N}(m, \sigma^2)$,

$$F_X(x) = \Phi\left(\frac{x - m}{\sigma}\right). \quad (10.5)$$

- (a) Here $n = 20$. So, we have $Y_n \sim \mathcal{N}(40 \times 20, 100 \times 20) = \mathcal{N}(800, 2000)$. For this random variable $m = 800$ and $\sigma = \sqrt{2000}$.

We want to find $P[Y_{20} > 1000]$ which is the same as $1 - P[Y_{20} \leq 1000]$. Expressing this quantity using cdf, we have

$$P[Y_{20} > 1000] = 1 - F_{Y_{20}}(1000).$$

Apply (10.5) to get

$$P[Y_{20} > 1000] = 1 - \Phi\left(\frac{1000 - 800}{\sqrt{2000}}\right) = 1 - \Phi(4.472) \approx Q(4.47) \approx \boxed{3.91 \times 10^{-6}}.$$

- (b) Here, the value of n is what we want. So, we will need to keep the formula in the general form. Again, from (10.5), for $Y_n \sim \mathcal{N}(40n, 100n)$, we have

$$P[Y_n > 1000] = 1 - F_{Y_n}(1000) = 1 - \Phi\left(\frac{1000 - 40n}{10\sqrt{n}}\right) = 1 - \Phi\left(\frac{100 - 4n}{\sqrt{n}}\right).$$

To find the value of n such that $P[Y_n > 1000] > 0.99$, we will first find the value of z which make

$$1 - \Phi(z) > 0.99. \quad (10.6)$$

At this point, we may try to solve for the value of Z by noting that (10.6) is the same as

$$\Phi(z) < 0.01. \quad (10.7)$$

Unfortunately, the tables that we have start with $\Phi(0) = 0.5$ and increase to something close to 1 when the argument of the Φ function is large. This means we can't directly find 0.01 in the table. Of course, 0.99 is in there and therefore we will need to solve (10.6) via another approach.

²Note that the expected value and the variance in this question are proportional to n . This naturally occurs when we consider the sum of i.i.d. random variables. The approximation by Gaussian random variable is a result of the central limit theorem (CLT).

To do this, we use another property of the Φ function. Recall that $1 - \Phi(z) = \Phi(-z)$. Therefore, (10.6) is the same as

$$\Phi(-z) > 0.99. \quad (10.8)$$

From our table, we can then conclude that (10.7) (which is the same as (10.8)) will happen when $-z > 2.33$. (If you have **MATLAB**, then you can get a more accurate answer of 2.3263.)

Now, plugging in $z = \frac{100-4n}{\sqrt{n}}$, we have $\frac{4n-100}{\sqrt{n}} > 2.33$. To solve for n , we first let $x = \sqrt{n}$. In which case, we have $\frac{4x^2-100}{x} > 2.33$ or, equivalently, $4x^2 - 2.33x - 100 > 0$. The two roots are $x = -4.717$ and $x > 5.3$. So, We need $x < -4.717$ or $x > 5.3$. Note that $x = \sqrt{n}$ and therefore can not be negative. So, we only have one case; that is, we need $x > 5.3$. Because $n = x^2$, we then conclude that we need $\boxed{n > 28.1 \text{ years}}$.